Domain selection for Gaussian Processes An application to electrocardiogram signals

Nicolás Hernández

University College London

Oct. 26th, 2023

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This work...



Gabriel Martos, PhD. Associate Professor at the Department of Mathematics and Statistics of Universidad Torcuato Di Tella, Buenos Aires, Argentina.



Hernández, N., & Martos, G. (2023). Domain Selection for Gaussian Process Data: An application to electrocardiogram signals. arXiv preprint arXiv:2306.00538.

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what is the deal with having too much data?



Overfitting

Estimation precision

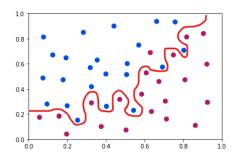
Computational complexity

Interpretability

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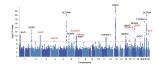
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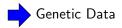
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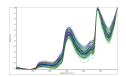


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We can find these issues in several applications











Health records, e.g. electrocardiograms

(some) solution to these problems...

How can we tackle these issues?



• Dimension reduction methods (e.g. PCA)

Regularisation (e.g. LASSO)

Variable selection

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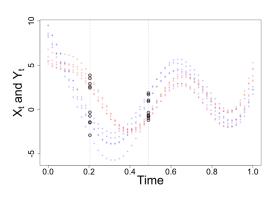
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Focus on variable selection...



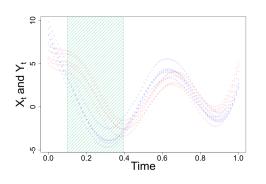
Select isolated variables

- find models with good prediction power,
- estimate the true "sparsity pattern".

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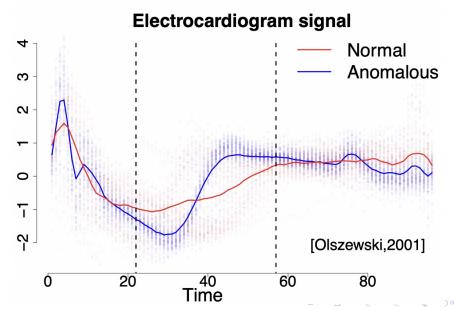
... but adding some restrictions: Domain Selection



Domain (interval) selection

- Select set of variables (intervals).
- Take advantage of the covariance structure.
- Variables are recorded almost continuously.

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For a pair of random processes, what is the region (domain) where they statistically differ the most?

- Inference:
 - Characterise and quantify the uncertainty around the estimation of the domain bounds.
 - Improve the power of a hypothesis test, i.e: two-sample test.
- **Prediction:** improve a classification model outcome, i.e. the misclassification error rate.
- Computational burden: reduce time and memory storage in the big data context. Future collection Strategy.

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compute differences with a local approach



Kullback-Leibler divergence



Let P and Q be two distributions of a continuous random variable x. Then the KL is:

$$\mathsf{KL}(P||Q) = \int_{-\infty}^{\infty} p(x) log(\frac{p(x)}{q(x)}) dx$$

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compute differences with a local approach



Kullback-Leibler divergence + Gaussian framework



$$KL(X||Y) = \frac{1}{2} \left(\frac{\sigma_X^2}{\sigma_Y^2} - 1 + \frac{(\mu_X - \mu_Y)^2}{\sigma_Y^2} + \ln\left(\frac{\sigma_Y^2}{\sigma_X^2}\right) \right).$$

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- Let $X(t) \sim GP(\mu_X(t), \sigma_X(t, s))$ be a GP indexed on the compact set $T \subset \mathbb{R}$, with:
 - $\mu_X(t) = E\{X(t)\}\$ and $\sigma_X(t,s) = E\{(X(t) \mu_X(t))(X(s) \mu_X(s))\}.$
 - Likewise for Y(t)

Learn the interval where X(t) and Y(t) statistically differ the most.

- Estimate a compact subset $A \subset T$ with $\lambda(A) > 0$:
 - (A) Differences in Mean. $t \in A$: $|\mu_X(t) \mu_Y(t)| > \nu$, and $t' \notin A$: $|\mu_X(t') \mu_Y(t')| \le \nu$.

Differences in Variance. $(t,s) \in A \times A$: $|\sigma_X(t,s) - \sigma_Y(t,s)| > \nu$, and $(t',s') \notin A \times A$: $|\sigma_X(t',s') - \sigma_Y(t',s')| \leq \nu$.

(A) and (B) simultaneously, possibly on different subsets A_{μ} and A_{σ} and for different ν_{μ} and ν_{σ} .

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 - (C) (A) and (B) simultaneously, possibly on different subsets A_{μ} and A_{σ} and for different ν_{μ} and ν_{σ} .

The KL divergence in the GP context

- In practice GP data is recorded over a finite grid $\mathcal{T}=(t_1,\ldots,t_p)$.
- In this case it corresponds to realisations of *p*-variate Gaussian random vectors and the KL has the following closed form:

$$\mathsf{KL}_{\mathcal{T}}(\mathbf{X}||\mathbf{Y}) \equiv \frac{1}{2} \left(\mathsf{tr} \left(\mathbf{\Sigma}_{\mathcal{T},Y}^{-1} \mathbf{\Sigma}_{\mathcal{T},X} - \mathbf{I}_{p} \right) + \Delta_{\mathcal{T}}^{\mathsf{T}} \mathbf{\Sigma}_{\mathcal{T},Y}^{-1} \Delta_{\mathcal{T}} + \ln \left(\frac{\det \mathbf{\Sigma}_{\mathcal{T},Y}}{\det \mathbf{\Sigma}_{\mathcal{T},X}} \right) \right),$$

where:

- $\bullet \ \Delta_{\mathcal{T}} = (\mu_{\mathcal{T},Y} \mu_{\mathcal{T},X}),$
- \bullet tr(Σ) denote the trace of Σ , and
- $det(\Sigma)$ the determinant of Σ .

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Domain Selection using the local-KL divergence

• For any subset $\mathcal{A} \subseteq \mathcal{T}$, we can also define the local–KL divergence as follows:

$$\mathsf{KL}_{\mathcal{A}}(X||Y) \equiv \frac{1}{2} \left(\mathsf{tr} \left(\mathbf{\Sigma}_{\mathcal{A},Y}^{-1} \mathbf{\Sigma}_{\mathcal{A},X} - \mathbf{I}_{|\mathcal{A}|} \right) + \Delta_{\mathcal{A}}^\mathsf{T} \mathbf{\Sigma}_{\mathcal{A},Y}^{-1} \Delta_{\mathcal{A}} + \mathsf{In} \left(\frac{\det \mathbf{\Sigma}_{\mathcal{A},Y}}{\det \mathbf{\Sigma}_{\mathcal{A},X}} \right) \right)$$

- ullet Consider $\mathcal{A} \in \mathcal{C}_{\mathcal{T}}$, collection of all contiguous subsets from \mathcal{T}
- $A^*(c)$, with $c \in (0,1)$, is the domain with maximum divergence

$$\max_{\mathcal{A} \in \mathcal{C}_{\mathcal{T}}} \mathsf{KL}_{\mathcal{A}}(X||Y), \text{ s.t. } \mathsf{len}(\mathcal{A}) \leq c\lambda(\mathcal{T}).$$

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Local-KL Divergence properties

The $KL_A(X||Y)$ is a set function that satisfies the following properties:

- (a) **Non-negative:** For fixed GPs X and Y, it holds: $\mathrm{KL}_{\mathcal{A}}(X||Y): \mathcal{P}_{\mathcal{T}} \to \mathbb{R}^+_0$ and $\mathrm{KL}_{\mathcal{A}}(X||Y) = 0$ if and only if $\mu_X(t) = \mu_Y(t)$ for all $t \in \mathcal{A}$ and $\sigma_X(t,s) = \sigma_Y(t,s)$ for all $(t,s) \in \mathcal{A} \times \mathcal{A}$.
- (b) The local KL divergence is **upper bounded** (i.e. $\mathrm{KL}_{\mathcal{T}}(X||Y) < \infty$) and a **monotone** set function (i.e. for $\mathcal{A}' \subseteq \mathcal{A}$ it holds: $\mathrm{KL}_{\mathcal{A}'}(X||Y) \le \mathrm{KL}_{\mathcal{A}}(X||Y)$)
- (c) The local divergence $\mathsf{KL}_{\mathcal{A}}(X||Y)$ is a **continuous** set function in $\mathcal{C}_{\mathcal{T}}$, the collection of all contiguous subsets from the ground set \mathcal{T} .

Properties $+ \mathcal{P}_{\mathcal{T}}$ is finite $\Rightarrow \mathcal{A}^*(c)$ exists.

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Local-KL divergence estimation

- Samples recorded over the same discrete grid \mathcal{T} : $\mathcal{D}_X = \{\mathbf{x}_i\}_{i=1}^n$ and $\mathcal{D}_Y = \{\mathbf{y}_j\}_{j=1}^m$ drawn from $GP(\mu_X(t), \sigma_X(t,s))$ and $GP(\mu_Y(t), \sigma_Y(t,s))$.
- The Maximum Likelihood estimates are given by:

$$\widehat{\boldsymbol{\mu}}_{\mathcal{T},X} = \frac{1}{n} \sum_{i=1}^{n} \mathbf{x}_{i}; \quad \widehat{\boldsymbol{\Sigma}}_{\mathcal{T},X} = \frac{1}{n} \sum_{i=1}^{n} (\mathbf{x}_{i} - \widehat{\boldsymbol{\mu}}_{\mathcal{T},X}) (\mathbf{x}_{i} - \widehat{\boldsymbol{\mu}}_{\mathcal{T},X})^{T},$$

(analogous expression holds for $\mu_{\mathcal{T},Y}, \Sigma_{\mathcal{T},Y}$).

• The trace and determinant are continuous functions in the space of real symmetric matrices, then $\widehat{\mathrm{KL}}_{\mathcal{A}}(X||Y)$ inherits (fixed p), consistency, asymptotic normality and efficiency,

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- Symmetrised KL divergence: $\frac{KL(X||Y) + KL(Y||X)}{2}$
- Regularised covariance matrix: $\widehat{\Sigma}_{\eta} = \eta \, \widehat{\Sigma} + (1 \eta) \operatorname{diag}(\widehat{\Sigma}), \text{ for } \eta \in [0, 1].$
- Sampling designs and misalignment: data asynchrony may act as a confounding factor when the aim is the estimation of the interval of local maximum KL divergence
- Smoothing: **low signal to noise ratio** leading to problems in the estimation of the interval of local maximum KL divergence

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One-shot experiment

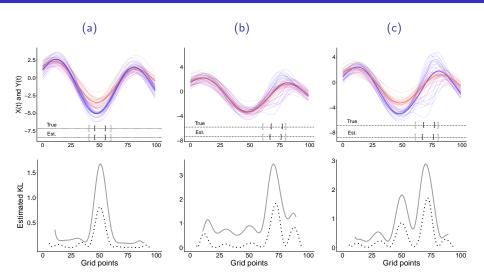


Figure: One shot experiment: Lower panels show the estimated local KL divergence for different interval lengths: c=0.1 (·····) and c=0.2 (—)

Monte Carlo experiment

- MC replicates = 1000
- Sample size: $n = \{50, 100, 500, 1000\}.$
- Grid points: $p = \{50, 100, 200, 500\}.$
- c possible lengths: (1% 99%).
- Av. Jaccard Distance: $AJD = E\left\{ \int_0^1 \left[1 \frac{|\mathcal{A}^*(c) \cap \widehat{\mathcal{A}}^*(c)|}{|\mathcal{A}^*(c) \cup \widehat{\mathcal{A}}^*(c)|} \right] \mathrm{d}c \right\}$

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Monte Carlo results

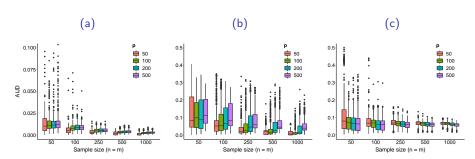


Figure: Empirical distribution of AIJD for different sample sizes n, m and grid resolution levels p.

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Computational complexity

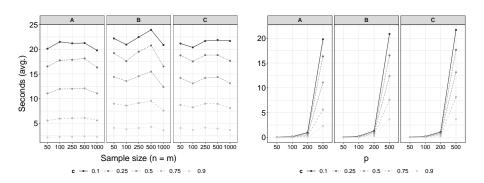


Figure: Average computational time required to estimate intervals of local maximum divergence.

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Monitoring Electrocardiogram signals

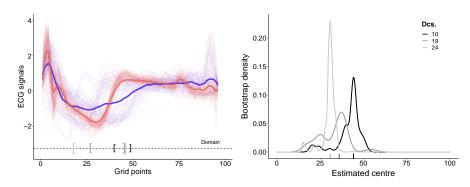


Figure: Left: ECG signals and selected domain for interval lengths 10,19 and 24 dcs. Right: Bootstrap densities for the interval centre (same lengths)

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ECG Classification

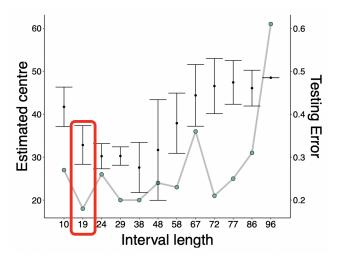


Figure: Estimated centre and classification error using a Discriminant model for different selected domains.

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- \checkmark Parameter of interest: interval \Rightarrow Covariance structure + almost continuously recorded data.
- ✓ Use the KL divergence and GP to develop an easy to implement algorithm for domain selection.
- \checkmark Propose an estimator for $\mathcal{A}^*(c)$, and a nonparametric approach to assess the estimation uncertainty
- \Rightarrow Consider other model frameworks (non Gaussian), e.g.: Wasserstein divergence.

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Some references

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Thank you for listening! Comments are very welcome.



(flash for arXiv pre-print)